

# Biostat3: Exercise on risk estimation

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Solutions are provided in Stata, R and SAS.

## Stata

In outline, we consider various approaches to estimate 5-year risks of colon cancer death for females following a colon cancer diagnosis by stage of disease.

We read in the dataset:

```
. set linesize 80
. cd /home/marcle/repos/biostat3_2014/tmp/
/home/marcle/repos/biostat3_2014/tmp
. use colon, clear
```

We then stset the data (after creating the id variable):

```
. capture drop id
. gen id = _n
. stset exit, fail(status==1) origin(dx) id(id) scale(365.24)

          id: id
failure event: status == 1
obs. time interval: (exit[_n-1], exit]
exit on or before: failure
t for analysis: (time-origin)/365.24
origin: time dx
```

```
-----
15564 total observations
      0 exclusions
-----
```

```
15564 observations remaining, representing
15564 subjects
 8369 failures in single-failure-per-subject data
58546.178 total analysis time at risk and under observation
                    at risk from t =          0
                    earliest observed entry t =          0
                    last observed exit t = 20.98073
```

As a first approach, we can calculate non-parametric risk (or failure) estimates from the Kaplan-Meier curves using the `sts list` command:

```
. sts list if agegrp==2 & stage!=0 & sex==2, by(stage) at(5 10) failure

failure _d: status == 1
analysis time _t: (exit-origin)/365.24
origin: time dx
id: id
```

Time	Beg. Total	Fail	Failure Function	Std. Error	[95% Conf. Int.]	
-----						
Localised						
5	793	310	0.2317	0.0117	0.2097	0.2556
10	338	52	0.2956	0.0138	0.2694	0.3237
Regional						
5	136	225	0.5299	0.0250	0.4819	0.5796
10	50	18	0.6109	0.0274	0.5577	0.6646
Distant						
5	87	1102	0.9072	0.0088	0.8890	0.9235
10	39	15	0.9276	0.0084	0.9101	0.9429

Note: Failure function is calculated over full data and evaluated at indicated times; it is not calculated from aggregates shown at left.

As a second approach, we can fit a Poisson regression with constant rates:

```
. streg i.sex##i.stage if agegrp==2 & stage!=0, dist(exp) nolog base

      failure _d:  status == 1
analysis time _t:  (exit-origin)/365.24
      origin:  time dx
      id:  id
```

Exponential regression -- log relative-hazard form

No. of subjects =	5,735	Number of obs =	5,735
No. of failures =	3,087		
Time at risk =	22828.45526		
		LR chi2(5) =	4300.57
Log likelihood =	-8039.8071	Prob > chi2 =	0.0000

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
-----						
sex						
Male	1	(base)				
Female	.8576709	.065315	-2.02	0.044	.7387515	.9957332
stage						
Localised	1	(base)				
Regional	2.642185	.2632503	9.75	0.000	2.173477	3.21197
Distant	13.62924	.9005805	39.53	0.000	11.97365	15.51374
sex#stage						
Female #						
Regional	1.258471	.1618644	1.79	0.074	.9780514	1.61929
Female #						
Distant	1.070361	.0947776	0.77	0.443	.8998258	1.273215
_cons						
	.0463934	.0026307	-54.15	0.000	.0415135	.051847

We can obtain confidence intervals for the rates using the lincom command:

```
. lincom _cons+2.sex, eform

( 1)  [_t]2.sex + [_t]_cons = 0
```

_t	exp(b)	Std. Err.	z	P> z	[95% Conf. Interval]	
(1)	.0397903	.0020227	-63.43	0.000	.0360171	.0439588

Using the confidence intervals for the rate, we can calculate the five-year risk and its 95% confidence intervals:

```
. di 1-exp(-.0397903*5)
.18041036
. di 1-exp(-.0360171*5)
.1648012
. di 1-exp(-.0439588*5)
.19731587
```

Finally, we can use Cox regression:

```
. stcox i.sex##i.stage if agegrp==2 & stage!=0, nolog base

      failure _d:  status == 1
analysis time _t:  (exit-origin)/365.24
              origin:  time dx
              id:  id
```

Cox regression -- Breslow method for ties

```
No. of subjects =          5,735                Number of obs   =          5,735
No. of failures =          3,087
Time at risk    = 22828.45526
Log likelihood   = -23864.82                    LR chi2(5)         =          2775.22
                                                Prob > chi2        =          0.0000
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
sex						
Male	1	(base)				
Female	.9418098	.0717916	-0.79	0.432	.8111074	1.093574
stage						
Localised	1	(base)				
Regional	2.285848	.2278625	8.29	0.000	1.880166	2.779064
Distant	8.842267	.5936013	32.47	0.000	7.752121	10.08572
sex#stage						
Female #						
Regional	1.179394	.1517318	1.28	0.200	.9165369	1.517637
Female #						
Distant	1.053203	.0932987	0.59	0.558	.885335	1.2529

We can pull out  $S_0(5)$  using the predict function with the basesurv argument. One way to get the times is the following:

```
. preserve
. sort _t
. predict S0, basesurv
```

```
(9,829 missing values generated)
. keep if agegrp==2 & stage!=0 & _t<5
(11,506 observations deleted)
. list _t S0 if _n==_N
```

```

+-----+
|          _t          S0 |
+-----+
4058. | 4.9802869    .7434324 |
+-----+
. restore
```

We can combine the baseline survival with the hazard ratio to calculate the risk. The risk equation is  $Risk = 1 - S_0(t)^{HR}$ . Plugging in these estimates for females at five years:

```
. display 1-.74343243^.9418098
.24363054
```

Confidence interval estimation can be done with the use of the `survi` user-contributed command:

```
. survci if agegrp==2 & stage!=0, survival at(sex=2, stage=1) outfile(pred, rep
> lace)
(sex=2.00; stage=1.00)
```

```
file pred.dta saved
. preserve
. use pred, clear
(estimates adjusted for sex=2; stage=1)
. sort _t
. keep if _t<5
(1,677 observations deleted)
. gen _fail = 1-_surv
. gen _fail_lb = 1-_ub
. gen _fail_ub = 1-_lb
. list _t _fail _fail_lb _fail_ub if _n==_N
```

```

+-----+
|          _t          _fail  _fail_lb  _fail_ub |
+-----+
4058. | 4.9802869    .2436305    .2230176    .2658031 |
+-----+
. restore
```

## R

Initially, load the required packages and read in the data.

```
options(width=80)
require(foreign)
require(survival)
colon <- read.dta("/home/marcle/repos/biostat3_2014/tmp/colon.dta")
nrow(colon)

[1] 15564
```

For the Kaplan-Meier analysis, we initially calculate the time on study (`tmExit`), subset the dataset, calculate the Kaplan-Meier estimates, and then summarise the risks.

```
require(dplyr)
colon <- mutate(colon, tmExit=as.numeric(exit - dx)/365.24)
tmp <- filter(colon, agegrp=="60-74" & sex=="Female")
fit1 <- survfit(Surv(tmExit,status=="Dead: cancer")~stage, data=tmp)
with(summary(fit1, time=c(5,10)),
      data.frame(strata, time, fail=1-surv, lower=1-upper, upper=1-lower))
## summary(fit1, time=c(5,10))
```

	strata	time	fail	lower	upper
1	stage=Unknown	5	0.3892831	0.3417862	0.4333527
2	stage=Unknown	10	0.4169903	0.3673982	0.4626946
3	stage=Localised	5	0.2316504	0.2083740	0.2542424
4	stage=Localised	10	0.2955891	0.2679427	0.3221915
5	stage=Regional	5	0.5299454	0.4783378	0.5764476
6	stage=Regional	10	0.6109463	0.5534227	0.6610603
7	stage=Distant	5	0.9071782	0.8882195	0.9229214
8	stage=Distant	10	0.9276370	0.9092668	0.9422878

We now re-do this analysis using Poisson regression and assuming a constant rate across time.

```
## streg i.sex##i.stage if agegrp==2 & stage!=0, dist(exp) nolog base
tmp <- filter(colon, agegrp=="60-74" & stage!="Unknown")
fit2 <- glm(I(status=="Dead: cancer")~factor(sex)*factor(stage)+
offset(log(tmExit)), data=tmp, family=poisson)
summary(fit2)
```

Call:

```
glm(formula = I(status == "Dead: cancer") ~ factor(sex) * factor(stage) +
    offset(log(tmExit)), family = poisson, data = tmp)
```

Deviance Residuals:

Min	1Q	Median	3Q	Max
-4.8782	-0.7964	-0.1320	1.3201	3.2919

Coefficients:

	Estimate	Std. Error	z value	Pr(> z )
(Intercept)	-3.07060	0.05670	-54.151	<2e-16
factor(sex)Female	-0.15353	0.07615	-2.016	0.0438
factor(stage)Regional	0.97161	0.09963	9.752	<2e-16
factor(stage)Distant	2.61222	0.06608	39.533	<2e-16
factor(sex)Female:factor(stage)Regional	0.22990	0.12862	1.787	0.0739
factor(sex)Female:factor(stage)Distant	0.06800	0.08855	0.768	0.4425

```
(Intercept) ***
factor(sex)Female *
factor(stage)Regional ***
factor(stage)Distant ***
factor(sex)Female:factor(stage)Regional .
factor(sex)Female:factor(stage)Distant
```

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for poisson family taken to be 1)

```
Null deviance: 14206.2 on 5734 degrees of freedom
Residual deviance: 9905.6 on 5729 degrees of freedom
AIC: 16092
```

Number of Fisher Scoring iterations: 7

We use the `predict` function to calculate the confidence interval for the log-rates, then calculate the rates and the risks.

```
pred2 <- predict(fit2, type="link",
  newdata=data.frame(sex="Female",stage="Localised",tmExit=1),
  se.fit=TRUE)
rates <- with(pred2,
  exp(cbind(fit, lower=fit-1.96*se.fit, upper=fit+1.96*se.fit)))
1-exp(-5*rates)
```

```
      fit      lower      upper
1 0.1804104 0.1648007 0.1973163
```

Finally, we fit a Cox regression using Breslow's methods for ties for comparison with Stata.

```
tmp <- filter(colon, agegrp=="60-74" & stage!="Unknown")
fit3 <- coxph(Surv(tmExit, status=="Dead: cancer")~factor(sex)*factor(stage),
  data=tmp, ties="breslow")
summary(fit3)
```

Call:

```
coxph(formula = Surv(tmExit, status == "Dead: cancer") ~ factor(sex) *
  factor(stage), data = tmp, ties = "breslow")
```

n= 5735, number of events= 3087

	coef	exp(coef)	se(coef)	z
factor(sex)Female	-0.05995	0.94181	0.07623	-0.786
factor(stage)Regional	0.82674	2.28585	0.09968	8.294
factor(stage)Distant	2.17954	8.84227	0.06713	32.466
factor(sex)Female:factor(stage)Regional	0.16500	1.17939	0.12865	1.283
factor(sex)Female:factor(stage)Distant	0.05184	1.05320	0.08859	0.585

Pr(>|z|)

factor(sex)Female	0.432
factor(stage)Regional	<2e-16 ***
factor(stage)Distant	<2e-16 ***
factor(sex)Female:factor(stage)Regional	0.200
factor(sex)Female:factor(stage)Distant	0.558

---  
 Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

	exp(coef)	exp(-coef)	lower .95
factor(sex)Female	0.9418	1.0618	0.8111
factor(stage)Regional	2.2858	0.4375	1.8802
factor(stage)Distant	8.8423	0.1131	7.7521
factor(sex)Female:factor(stage)Regional	1.1794	0.8479	0.9165
factor(sex)Female:factor(stage)Distant	1.0532	0.9495	0.8853

upper .95

factor(sex)Female	1.094
factor(stage)Regional	2.779
factor(stage)Distant	10.086
factor(sex)Female:factor(stage)Regional	1.518
factor(sex)Female:factor(stage)Distant	1.253

```
Concordance= 0.753 (se = 0.005 )
Rsquare= 0.384 (max possible= 1 )
Likelihood ratio test= 2775 on 5 df, p=0
Wald test = 2480 on 5 df, p=0
Score (logrank) test = 3259 on 5 df, p=0
```

To calculate the predicted risks, we `survfit` the `coxph` model for given covariates, use `summary` for the specific times, and calculate the risks. Interestingly, these estimates are not identical to Stata, although the Cox fit is very similar.

```
pred3 <- survfit(fit3, newdata=data.frame(sex="Female",stage="Localised"))
with(summary(pred3, time=c(5,10)),
      data.frame(time, fail=1-surv, lower=1-upper, upper=1-lower))
```

	time	fail	lower	upper
1	5	0.2420369	0.2203020	0.2631659
2	10	0.2881452	0.2630863	0.3123519

## SAS

Initially, set the environment and read in the data. The `%reset` macro is specific to my set-up, so this can be ignored.

```
1982 options linesize=80;
1983 %macro reset;
1984     options obs=max nosyntaxcheck replace;
1985 %mend reset;
1986 %reset;
NOTE: The PROCEDURE SQL printed page 303.
NOTE: PROCEDURE SQL used (Total process time):
      real time          13.42 seconds
      cpu time           0.01 seconds

1987 proc import datafile="/home/marcle/repos/biostat3_2014/tmp/colon.dta"
1987! out=colon replace;
1988 run;
NOTE: The import data set has 15564 observations and 13 variables.
NOTE: WORK.COLON data set was successfully created.
NOTE: PROCEDURE IMPORT used (Total process time):
      real time          0.00 seconds
      cpu time           0.01 seconds
```

Now, calculate the follow-up time and calculate the Kaplan-Meier estimators. I like to use SQL for data manipulation – one could do this differently using data steps and merging.

```
1990 data colon;
1991     set colon;
1992     tmexit = (exit-dx)/365.24;
1993 run;
NOTE: There were 15564 observations read from the data set WORK.COLON.
NOTE: The data set WORK.COLON has 15564 observations and 14 variables.
NOTE: DATA statement used (Total process time):
      real time          0.00 seconds
      cpu time           0.00 seconds

1994 proc lifetest data=colon notable outsurv=surv;
1995     time tmexit*status(0,2,4);
1996     strata stage;
1997     where agegrp=2 and sex=2;
1998 run;
NOTE: The LOGLOG transform is used to compute the confidence limits for the
survivor function. To suppress using this transform, specify
CONFTYPE=LINEAR in the PROC LIFETEST statement.
```

The LIFETEST Procedure

Summary of the Number of Censored and Uncensored Values

Stratum	stage	Total	Failed	Censored	Percent Censored
1	Distant	1254	1123	131	10.45
2	Localised	1488	387	1101	73.99
3	Regional	465	248	217	46.67
4	Unknown	487	190	297	60.99
-----					
Total		3694	1948	1746	47.27
					305
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The LIFETEST Procedure

Testing Homogeneity of Survival Curves for tmexit over Strata

Rank Statistics

stage	Log-Rank	Wilcoxon
Distant	735.62	2045502
Localised	-630.44	-1706540
Regional	-11.76	-124435
Unknown	-93.42	-214527

Covariance Matrix for the Log-Rank Statistics

stage	Distant	Localised	Regional	Unknown
Distant	291.196	-185.943	-51.959	-53.294
Localised	-185.943	468.352	-133.362	-149.047
Regional	-51.959	-133.362	222.430	-37.109
Unknown	-53.294	-149.047	-37.109	239.451

Covariance Matrix for the Wilcoxon Statistics

stage	Distant	Localised	Regional	Unknown
Distant	2.4668E9	-1.549E9	-4.606E8	-4.57E8
Localised	-1.549E9	3.3774E9	-9.07E8	-9.211E8
Regional	-4.606E8	-9.07E8	1.6281E9	-2.604E8
Unknown	-4.57E8	-9.211E8	-2.604E8	1.6385E9

Test of Equality over Strata

Test	Chi-Square	DF	Pr > Chi-Square
Log-Rank	1945.9735	3	<.0001



Wilcoxon 1772.5785 3 <.0001  
 -2Log(LR) 2530.4314 3 <.0001

NOTE: The data set WORK.SURV has 2380 observations and 7 variables.

NOTE: The PROCEDURE LIFETEST printed pages 304-305.

NOTE: PROCEDURE LIFETEST used (Total process time):

real time 0.00 seconds  
 cpu time 0.00 seconds

```
1999 proc sql noerrorstop;
2000     select distinct stage, tmexit, 1-survival as failure, 1-sdf_ucl as
2000! fail_lcl, 1-sdf_lcl as fail_ucl
2001     from surv
2002     natural join (select stage, max(tmexit) as tmexit from surv where
2002! tmexit<5 and _censor_=0 group by stage)
2003     where _censor_=0
2004     order by stage;
```

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stage	tmexit	failure	fail_lcl	fail_ucl
0	4.963859	0.389283	0.345338	0.436735
1	4.958384	0.23165	0.209662	0.255551
2	4.709232	0.529945	0.481908	0.579625
3	4.87077	0.907178	0.889008	0.923494

```
2005 quit;
```

NOTE: The PROCEDURE SQL printed page 306.

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NOTE: PROCEDURE SQL used (Total process time):

real time 0.00 seconds  
 cpu time 0.01 seconds

Now, fit the Poisson regression.

```
2007 data colon;
2008     set colon;
2009     event = (status=1);
2010     logpt=log(tmExit);
2011 run;
```

NOTE: There were 15564 observations read from the data set WORK.COLON.

NOTE: The data set WORK.COLON has 15564 observations and 16 variables.

NOTE: DATA statement used (Total process time):

real time 0.00 seconds  
 cpu time 0.01 seconds

```
2012 ods output estimates=est;
2013 proc genmod data=colon;
2014     class sex stage(ref='Localised');
2015     model event = sex|stage / dist=poisson offset=logpt;
2016     estimate "Rate"
2017         intercept 1
2018         sex 1 0
2019         stage 0 0 1
2020         sex*stage 0 0 1 0 0 0;
2021     where agegrp=2 and stage~=0;
2022 run;
```

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The GENMOD Procedure

Model Information

Data Set	WORK.COLON
Distribution	Poisson
Link Function	Log
Dependent Variable	event
Offset Variable	logpt

Number of Observations Read	5735
Number of Observations Used	5735

Class Level Information

Class	Levels	Values
sex	2	Female Male
stage	3	Distant Regional Localised

Parameter Information

Parameter	Effect	sex	stage
Prm1	Intercept		
Prm2	sex	Female	
Prm3	sex	Male	
Prm4	stage		Distant
Prm5	stage		Regional
Prm6	stage		Localised
Prm7	sex*stage	Female	Distant
Prm8	sex*stage	Female	Regional
Prm9	sex*stage	Female	Localised
Prm10	sex*stage	Male	Distant
Prm11	sex*stage	Male	Regional
Prm12	sex*stage	Male	Localised

Criteria For Assessing Goodness Of Fit

Criterion	DF	Value	Value/DF
Deviance	5729	9905.6142	1.7290
Scaled Deviance	5729	9905.6142	1.7290
Pearson Chi-Square	5729	52631.0454	9.1868
Scaled Pearson X2	5729	52631.0454	9.1868
Log Likelihood		-8039.8071	
Full Log Likelihood		-8039.8071	
AIC (smaller is better)		16091.6142	
AICC (smaller is better)		16091.6289	
BIC (smaller is better)		16131.5403	

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The GENMOD Procedure

Algorithm converged.  
NOTE: Algorithm converged.

Analysis Of Maximum Likelihood Parameter Estimates

Parameter	DF	Estimate	Standard Error	Wald 95% Confidence Limits
Intercept	1	-3.0706	0.0567	-3.1817 -2.9595
sex Female	1	-0.1535	0.0762	-0.3028 -0.0043
sex Male	0	0.0000	0.0000	0.0000 0.0000
stage Distant	1	2.6122	0.0661	2.4827 2.7417
stage Regional	1	0.9716	0.0996	0.7763 1.1669
stage Localised	0	0.0000	0.0000	0.0000 0.0000
sex*stage Female Distant	1	0.0680	0.0885	-0.1056 0.2415
sex*stage Female Regional	1	0.2299	0.1286	-0.0222 0.4820
sex*stage Female Localised	0	0.0000	0.0000	0.0000 0.0000
sex*stage Male Distant	0	0.0000	0.0000	0.0000 0.0000
sex*stage Male Regional	0	0.0000	0.0000	0.0000 0.0000
sex*stage Male Localised	0	0.0000	0.0000	0.0000 0.0000
Scale	0	1.0000	0.0000	1.0000 1.0000

Analysis Of Maximum Likelihood Parameter Estimates

Parameter	Wald Chi-Square	Pr > ChiSq
Intercept	2932.28	<.0001
sex Female	4.06	0.0438
sex Male	.	.
stage Distant	1562.85	<.0001
stage Regional	95.10	<.0001
stage Localised	.	.
sex*stage Female Distant	0.59	0.4425
sex*stage Female Regional	3.19	0.0739
sex*stage Female Localised	.	.
sex*stage Male Distant	.	.
sex*stage Male Regional	.	.
sex*stage Male Localised	.	.
Scale		

NOTE: The scale parameter was held fixed.  
NOTE: The scale parameter was held fixed.

Contrast Estimate Results

Label	Mean Estimate	Mean Confidence Limits	L'Beta Estimate	Standard Error	Alpha
Rate	0.0398	0.0360 0.0440	-3.2241	0.0508	0.05

The GENMOD Procedure

Contrast Estimate Results

Label	L'Beta Confidence Limits	Chi- Square	Pr > ChiSq
Rate	-3.3238      -3.1245	4022.9	<.0001

NOTE: The data set WORK.EST has 1 observations and 11 variables.

NOTE: The PROCEDURE GENMOD printed pages 307-309.

NOTE: PROCEDURE GENMOD used (Total process time):

real time            0.02 seconds  
cpu time             0.02 seconds

2023 proc sql noerrorstop;

2024        select label, 1-exp(-MeanEstimate\*5) as risk, 1-exp(-MeanLowerCL\*5) as

2024! lower, 1-exp(-MeanUpperCL\*5) as upper from est;

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Label	risk	lower	upper
Rate	0.18041	0.164801	0.197316

2025 quit;

NOTE: The PROCEDURE SQL printed page 310.

NOTE: PROCEDURE SQL used (Total process time):

real time            0.00 seconds  
cpu time             0.00 seconds

Finally, fit the Cox regression.

2027 data colon\_subset;

2028        sex=2; agegrp=2;

2029        do stage=1 to 3;

2030            output;

2031            end;

2032        format sex sexa. stage stagea. agegrp agegrpa.;

2033 run;

NOTE: The data set WORK.COLON\_SUBSET has 3 observations and 3 variables.

NOTE: DATA statement used (Total process time):

real time            0.00 seconds  
cpu time             0.01 seconds

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2034 proc phreg data=colon;

2035        class sex stage(ref='Localised');

2036        model tmExit\*status(0,2,4)=sex|stage;

2037        baseline covariates=colon\_subset out=pred3 survival=\_all\_;

2038        where agegrp=2 and stage~=0;

2039 run;

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The PHREG Procedure

Model Information

Data Set

WORK.COLON

```

Dependent Variable      tmexit
Censoring Variable     status      Vital status at last contact
Censoring Value(s)    0 2 4
Ties Handling          Breslow

```

```

Number of Observations Read      5735
Number of Observations Used     5735

```

Class Level Information

Class	Value	Design Variables	
sex	Female	1	
	Male	0	
stage	Distant	1	0
	Localised	0	0
	Regional	0	1

Summary of the Number of Event and Censored Values

Total	Event	Censored	Percent Censored
5735	3087	2648	46.17

Convergence Status

Convergence criterion (GCONV=1E-8) satisfied.

NOTE: Convergence criterion (GCONV=1E-8) satisfied.

Model Fit Statistics

Criterion	Without Covariates	With Covariates
-2 LOG L	50504.854	47729.639
AIC	50504.854	47739.639
SBC	50504.854	47769.814

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The PHREG Procedure

Testing Global Null Hypothesis: BETA=0

Test	Chi-Square	DF	Pr > ChiSq
Likelihood Ratio	2775.2151	5	<.0001
Score	3259.1938	5	<.0001

Wald 2479.6926 5 <.0001

Type 3 Tests

Effect	DF	Wald Chi-Square	Pr > ChiSq
sex	1	0.6186	0.4316
stage	2	1126.6664	<.0001
sex*stage	2	1.6516	0.4379

Analysis of Maximum Likelihood Estimates

Parameter	DF	Parameter Estimate	Standard Error	Chi-Square	Pr > ChiSq
sex Female	1	-0.05995	0.07623	0.6186	0.4316
stage Distant	1	2.17954	0.06713	1054.0599	<.0001
stage Regional	1	0.82678	0.09968	68.7919	<.0001
sex*stage Female Distant	1	0.05184	0.08859	0.3424	0.5584
sex*stage Female Regional	1	0.16502	0.12865	1.6453	0.1996

Analysis of Maximum Likelihood Estimates

Parameter	Hazard Ratio
sex Female	.
stage Distant	.
stage Regional	.
sex*stage Female Distant	.
sex*stage Female Regional	.

Analysis of Maximum Likelihood Estimates

Parameter	Label
sex Female	Sex Female
stage Distant	Clinical stage at diagnosis Distant
stage Regional	Clinical stage at diagnosis Regional
sex*stage Female Distant	Sex Female * Clinical stage at diagnosis Distant
sex*stage Female Regional	Sex Female * Clinical stage at diagnosis Regional

NOTE: Hazard ratios that cannot be conveniently calculated or displayed are set to missing in the ParameterEstimates table. Use the HAZARDRATIO statement to compute the needed hazard ratios.

NOTE: The data set WORK.PRED3 has 1065 observations and 8 variables.

NOTE: The PROCEDURE PHREG printed pages 311-312.

NOTE: PROCEDURE PHREG used (Total process time):

real time 0.01 seconds  
cpu time 0.01 seconds

```

2040 proc sql noerrorstop;
2041     select distinct stage, tmexit, 1-survival as failure, 1-UpperSurvival
2041! as fail_lcl, 1-LowerSurvival as fail_ucl
2042     from pred3
2043     natural join (select stage, max(tmexit) as tmexit from pred3 where

```

```
2043! tmexit<5 group by stage)
2044     where sex=2 and agegrp=2
2045     order by stage;
```

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stage	tmexit	failure	fail_lcl	fail_ucl
1	4.958384	0.242035	0.2203	0.263164
2	4.958384	0.526277	0.47947	0.568875
3	4.958384	0.92428	0.910237	0.936127

I always forget that SAS output is so verbose. Questions or revisions: please direct to me ([mark.clements@ki.se](mailto:mark.clements@ki.se)).