

# BIOSTAT III: Survival Analysis for Epidemiologists in Stata: Take-home examination

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## Instructions

- **The examination is individual-based: you are not allowed to cooperate with anyone**, although you are encouraged to consult the available literature. The examiner will use Urkund in order to assess potential plagiarism.
- The examination will be made available by noon on Wednesday 19 February 2020 and **the examination is due by 17:00 on Wednesday 26 February 2020**.
- The examination will be graded and results returned to you by Wednesday 4 March 2020.
- The examination is in two parts. You need to score at least 8/15 for Part 1 focused on rates and general regression modelling and 13/24 in Part 2 on survival analysis to pass the examination.
- Do not write answers by hand: please use Word, L<sup>A</sup>T<sub>E</sub>X or a similar format for your examination report and submit the report **as a PDF file**.
- Motivate all answers in your examination report. Define any notation that you use for equations. The examination report should be written in English.
- Email the examination report containing the answers **as a PDF file** to [gunilla.nilsson.roos@ki.se](mailto:gunilla.nilsson.roos@ki.se). **Write your name in the email, but do NOT write your name or otherwise reveal your identity in the document containing the answers.**

## Part 1

The **DMepi2** dataset includes simulated data on all cause mortality rates for those with and without diabetes in Denmark for 1996–2015. The dataset has the following columns:

**sex** a factor with levels 1=M, 2=F

**A** One-year age class, 0–99 years

**P** Calendar year, 1996–2016

**diab** Indicator for persons with diabetes (1=yes, 0=no)

**Y** Person-years

**D** Number of deaths

**R** Rates (=D/Y)

Q1

- (a) The age-specific mortality rates by sex and diabetes status for 2016 are shown in Figure 1. Carefully describe the pattern of rates by age, sex and diabetes status. (2 pts)

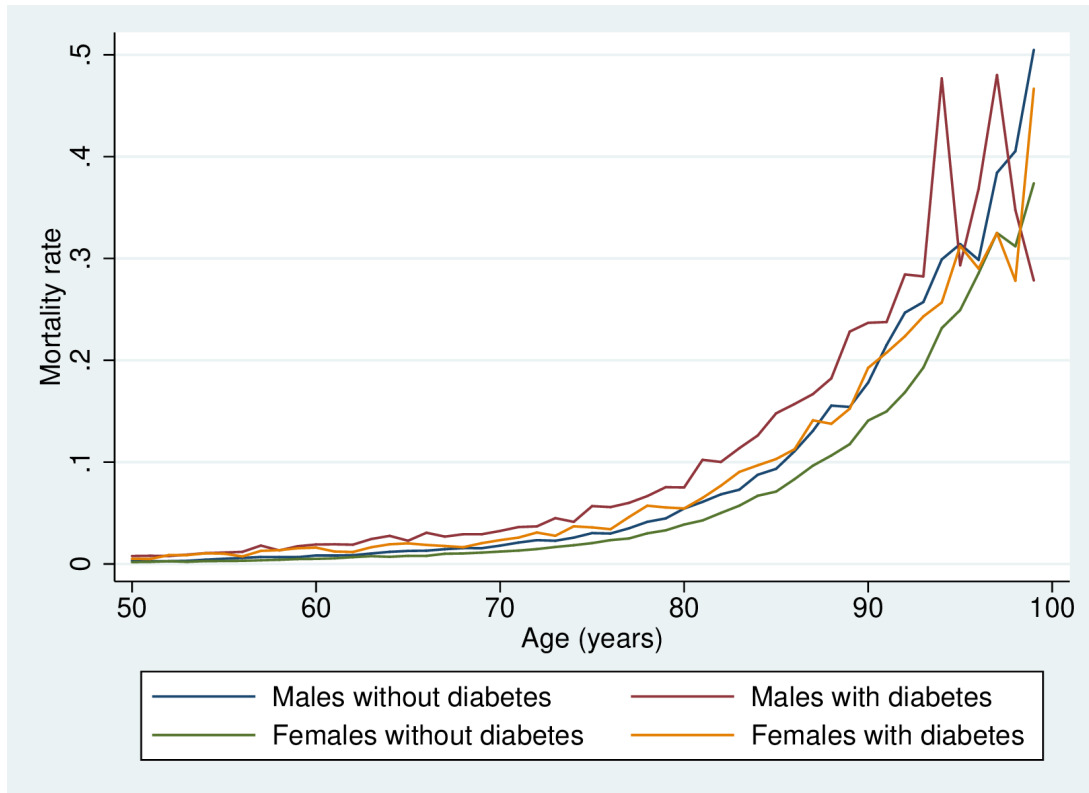


Figure 1: Age-specific mortality rates by sex and presence or absence of diabetes, Denmark 2016.

The following code and output is used to model the mortality rates by diabetes status for males and females separately for the 2016 calendar year:

```
use DMepi2, clear
keep if P==2016
poisson D A diab if sex==1, exp(Y)
```

(7,997 observations deleted)

```
Poisson regression                    Number of obs   =          199
LR chi2(2)                          =       72116.18
Prob > chi2                          =          0.0000
Pseudo R2                            =          0.9747

Log likelihood = -936.90877
```

| D     | Coef.     | Std. Err.  | z       | P> z  | [95% Conf. Interval] |
|-------|-----------|------------|---------|-------|----------------------|
| A     | .0989664  | .0004939   | 200.36  | 0.000 | .0979983 .0999345    |
| diab  | .4968066  | .0150224   | 33.07   | 0.000 | .4673633 .5262499    |
| _cons | -10.80233 | .0372321   | -290.13 | 0.000 | -10.8753 -10.72936   |
| ln(Y) | 1         | (exposure) |         |       |                      |

```
poisson D A diab if sex==2, exp(Y)
```

```

Poisson regression                                Number of obs   =      199
                                                LR chi2(2)      =    80738.73
                                                Prob > chi2     =      0.0000
Log likelihood = -873.46624                    Pseudo R2       =      0.9788

```

| D     | Coef.     | Std. Err.  | z       | P> z  | [95% Conf. Interval] |
|-------|-----------|------------|---------|-------|----------------------|
| A     | .1076621  | .0005173   | 208.11  | 0.000 | .1066481 .108676     |
| diab  | .451282   | .0164922   | 27.36   | 0.000 | .4189578 .4836061    |
| _cons | -11.78155 | .0415506   | -283.55 | 0.000 | -11.86299 -11.70011  |
| ln(Y) | 1         | (exposure) |         |       |                      |

- (b) Write out the regression model for males. As a reminder, please explain all of your notation. (2 pts)
- (c) What are the mortality rate ratios and 95% confidence intervals for those with diabetes compared with those without diabetes for (i) males and (ii) females? (2 pts)

The following interaction model and linear combination can be used to compare the mortality rate ratio of diabetes for males with the mortality rate ratio of diabetes for females. As a reminder, `baselevels` adds the base or reference level for a factor variable to the output.

```

poisson D A diab##sex, exp(Y) baselevels
lincom 1.diab + 1.diab#2.sex

```

```

Poisson regression                                Number of obs   =      398
                                                LR chi2(4)      =   152711.67
                                                Prob > chi2     =      0.0000
Log likelihood = -1884.4039                    Pseudo R2       =      0.9759

```

| D        | Coef.     | Std. Err.  | z       | P> z  | [95% Conf. Interval] |
|----------|-----------|------------|---------|-------|----------------------|
| A        | .103223   | .0003568   | 289.26  | 0.000 | .1025236 .1039224    |
| diab     |           |            |         |       |                      |
| 0        | 0         | (base)     |         |       |                      |
| 1        | .4854449  | .0149829   | 32.40   | 0.000 | .4560789 .5148108    |
| sex      |           |            |         |       |                      |
| M        | 0         | (base)     |         |       |                      |
| F        | -.3127235 | .0100123   | -31.23  | 0.000 | -.3323473 -.2930997  |
| diab#sex |           |            |         |       |                      |
| 1#F      | -.0245248 | .0222395   | -1.10   | 0.270 | -.0681135 .0190639   |
| _cons    | -11.11888 | .0275994   | -402.87 | 0.000 | -11.17298 -11.06479  |
| ln(Y)    | 1         | (exposure) |         |       |                      |

( 1) [D]1.diab + [D]1.diab#2.sex = 0

| D | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] |
|---|-------|-----------|---|------|----------------------|
|---|-------|-----------|---|------|----------------------|

```
(1) |      .4609201      .0164799      27.97      0.000      .4286201      .4932201
```

- 
- (d) Write out the regression equation for the interaction model. (1pt)
- (e) What are the mortality rate ratios and 95% confidence intervals for those with diabetes compared with those without diabetes for (i) males and (ii) females? Why are these estimates different to the estimates in (c)? (2pts)
- (f) Formally test for whether the two mortality rate ratios for males and females in (e) are different. Explain how you undertook the test and interpret the findings. (2pts)

We now model calendar period as a continuous, linear effect using a main effects model:

```
use DMepi2, clear
poisson D A sex diab P, exp(Y) baselevels
```

```
Poisson regression              Number of obs      =          8,395
                               LR chi2(4)              =    3398186.17
                               Prob > chi2              =          0.0000
Log likelihood = -37872.375     Pseudo R2         =          0.9782
```

```
-----
          D |      Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
          A |   .0978487   .0000704   1390.13  0.000   .0977108   .0979867
        sex |  -.3675351   .0018946   -194.00  0.000  -.3712484  -.3638218
        diab |   .5393626   .0026551   203.14  0.000   .5341586   .5445665
          P |  -.0253965   .0001544   -164.52  0.000  -.0256999  -.0250939
        _cons |  40.87618   .3096228   132.02  0.000   40.26933   41.48303
ln(Y) |              1 (exposure)
```

- 
- (g) Write out the regression equation for this model. (1pt)
- (h) How would you interpret the parameter P and its 95% confidence interval? (1pt)
- (i) How would you interpret the parameter `_cons`? Why is the value so large? (2pts)

## Part 2

### Q2

We now use data from the German Breast Cancer Study Group (GBCSG) on a randomised study of hormonal treatment and the duration of chemotherapy in node-positive breast cancer patients (see <https://doi.org/10.1200/JCO.1994.12.10.2086>). The event considered was time to recurrence of breast cancer or death due to breast cancer ("recurrence-free survival"). The main study found no effect associated with the duration of chemotherapy on recurrence-free survival. The code-book for some of the dataset is shown below:

```
use brcancer, clear
keep id hormon x1 x6 rectime censrec
egen x1cat = cut(x1), at(0,45,60,80) label
egen x6cat = cut(x6), at(0,20,2380) label
codebook
```

(German breast cancer data)  
 (2 missing values generated)  
 (1 missing value generated)

-----  
 id Individual ID  
 -----

type: numeric (int)  
 range: [1,686] units: 1  
 unique values: 686 missing .: 0/686  
 mean: 343.5  
 std. dev: 198.175  
 percentiles: 10% 25% 50% 75% 90%  
 69 172 343.5 515 618

-----  
 hormon hormonal therapy  
 -----

type: numeric (byte)  
 label: hormon  
 range: [0,1] units: 1  
 unique values: 2 missing .: 0/686  
 tabulation: Freq. Numeric Label  
 440 0 Standard treatment  
 246 1 Hormonal treatment

-----  
 x1 age, years  
 -----

type: numeric (byte)  
 range: [21,80] units: 1  
 unique values: 54 missing .: 0/686  
 mean: 53.0525  
 std. dev: 10.1207  
 percentiles: 10% 25% 50% 75% 90%  
 40 46 53 61 65

-----  
 x6 progesterone receptor, fmol  
 -----

type: numeric (int)  
 range: [0,2380] units: 1  
 unique values: 242 missing .: 0/686  
 mean: 109.996  
 std. dev: 202.332  
 percentiles: 10% 25% 50% 75% 90%  
 0 7 32.5 132 312

-----  
 rectime recurrence-free survival time, days  
 -----

type: numeric (int)  
 range: [8,2659] units: 1  
 unique values: 574 missing .: 0/686  
 mean: 1124.49  
 std. dev: 642.792

```

percentiles:      10%      25%      50%      75%      90%
                  322      567      1084     1685     2014

```

```
-----
censrec                                                    censoring indicator
-----
```

```

      type: numeric (byte)
      label: event
      range: [0,1]
unique values: 2
      tabulation: Freq.  Numeric  Label
                  387      0  censored
                  299      1  event
                units: 1
                missing .: 0/686

```

```
-----
x1cat                                                    (unlabeled)
-----
```

```

      type: numeric (float)
      label: x1cat
      range: [0,2]
unique values: 3
      tabulation: Freq.  Numeric  Label
                  131      0  0-
                  344      1  45-
                  209      2  60-
                   2      .
                units: 1
                missing .: 2/686

```

```
-----
x6cat                                                    (unlabeled)
-----
```

```

      type: numeric (float)
      label: x6cat
      range: [0,1]
unique values: 2
      tabulation: Freq.  Numeric  Label
                  269      0  0-
                  416      1  20-
                   1      .
                units: 1
                missing .: 1/686

```

- (a) For this dataset, we have the time from randomisation to recurrence or breast cancer death. Assume we also had (i) the date of birth, (ii) date of cancer diagnosis, (iii) date of randomisation and (iv) date of recurrence or death. Discuss which time scales you could use for your analysis, describing their advantages and disadvantages. (2pts)

We now `stset` for the time from randomisation to time of recurrence or death – that is, we are modelling for recurrence-free survival. There were 299 events and the event times are in days from randomisation.

- (b) The Kaplan-Meier estimators for the survival functions by hormonal treatment are shown in Figure 2. Carefully describe and interpret the two survival curves. (2pts)

```

stset rectime, fail(censrec==1)
sts graph, by(hormon) title("")

      failure event:  censrec == 1
obs. time interval:  (0, rectime]

```

```
exit on or before: failure
```

```
-----  
686 total observations  
0 exclusions  
-----
```

```
686 observations remaining, representing  
299 failures in single-record/single-failure data  
771,400 total analysis time at risk and under observation  
at risk from t = 0  
earliest observed entry t = 0  
last observed exit t = 2,659  
  
failure _d: censrec == 1  
analysis time _t: rectime
```

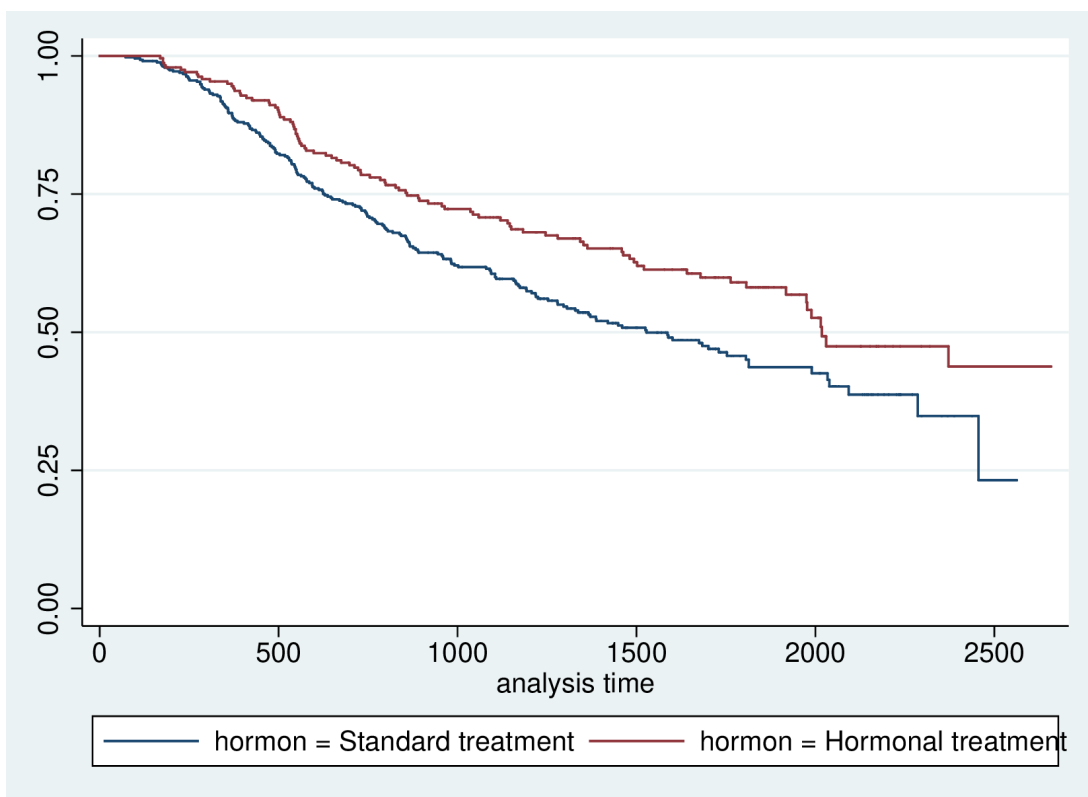


Figure 2: Kaplan-Meier survival curves by hormonal treatment, German Breast Cancer Study Group

(c) For the following log-rank test, state the null hypothesis and interpret the test. (1pt)

```
sts test hormon
```

```
failure _d: censrec == 1  
analysis time _t: rectime  
Log-rank test for equality of survivor functions
```

```
-----  
hormon      | Events      Events  
            | observed    expected  
-----+-----
```

|                    |  |           |        |
|--------------------|--|-----------|--------|
| Standard treatment |  | 205       | 180.34 |
| Hormonal treatment |  | 94        | 118.66 |
| -----+-----        |  |           |        |
| Total              |  | 299       | 299.00 |
|                    |  | chi2(1) = | 8.56   |
|                    |  | Pr>chi2 = | 0.0034 |

(d.i) Write out the regression equation for the Cox model specified in the following code and output. (2pts)

(d.ii) Based on the following output, discuss whether there is any evidence that hormonal treatment is associated with recurrence-free survival. (2pts)

```
stcox i.x1cat i.x6cat hormon, baselevels nohr

      failure _d:  censrec == 1
      analysis time _t:  rectime
Refining estimates:
Cox regression -- Breslow method for ties
No. of subjects =          684          Number of obs   =          684
No. of failures =          298
Time at risk    =          770171
Log likelihood  = -1753.6988          LR chi2(4)       =          55.37
                                          Prob > chi2     =          0.0000
-----+-----
      _t |          Coef.   Std. Err.      z    P>|z|      [95% Conf. Interval]
-----+-----
      x1cat |
      0- |              0   (base)
      45- |   -.3036346   .1501781    -2.02   0.043   -.5979782   -.0092909
      60- |   -.1439799   .1642114    -0.88   0.381   -.4658284   .1778685
      |
      x6cat |
      0- |              0   (base)
      20- |   -.7635767   .1164474    -6.56   0.000   -.9918095   -.5353439
      |
      hormon |   -.3453868   .1274155    -2.71   0.007   -.5951166   -.0956569
-----+-----
```

(e) Based on the following Schoenfeld residuals table, is there any evidence for non-proportionality in the modelled covariates? Interpret the table and explain your reasoning. (2pt)

```
estat phtest, detail

Test of proportional-hazards assumption
Time: Time
-----+-----
      |          rho          chi2          df          Prob>chi2
-----+-----
0b.x1cat |          .          .          1          .
1.x1cat |   0.06832          1.38          1          0.2399
2.x1cat |   0.09127          2.44          1          0.1179
0b.x6cat |          .          .          1          .
-----+-----
```



|             |  |         |      |   |        |
|-------------|--|---------|------|---|--------|
| 1.x6cat     |  | 0.11709 | 4.01 | 1 | 0.0453 |
| hormon      |  | 0.00780 | 0.02 | 1 | 0.8933 |
| -----+      |  |         |      |   |        |
| global test |  |         | 7.01 | 4 | 0.1351 |
| -----       |  |         |      |   |        |

(f) Based on the previous table and the following plot (Figure 3), how would you expect the hazard ratio for progesterone receptor to vary by time since randomisation? Explain your reasoning. (2pts)

estat phtest, plot(1.x6cat)

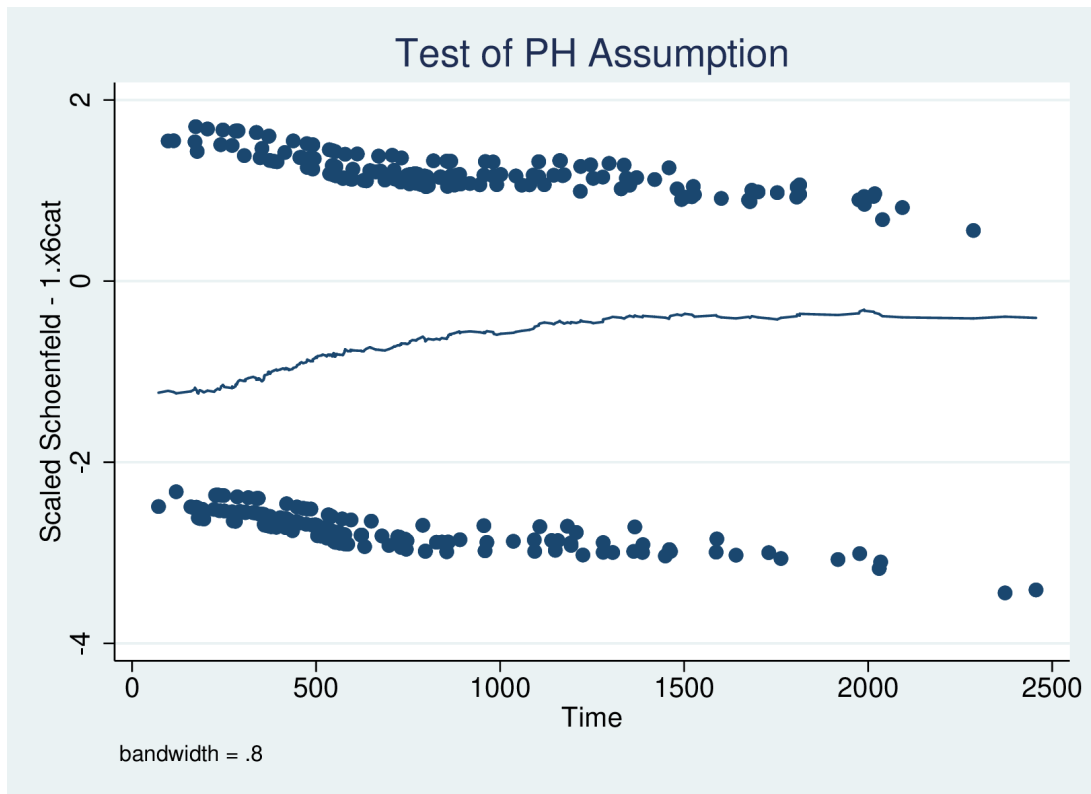


Figure 3: Schoenfeld residual plot for progesterone receptor  $\geq 20$  fmol, German Breast Cancer Study Group

(g) We now fit a flexible parametric survival model adjusting for **x1cat**, **x6cat** and **hormon** (see the following output). How is this model different to the model in (d)? (2pts)

```
stpm2 i.x1cat i.x6cat hormon, baselevels df(4) scale(haz)
est store main_effects
```

```
Log likelihood = -643.9481          Number of obs   =      684
-----+-----
          |      Coef.   Std. Err.      z    P>|z|      [95% Conf. Interval]
-----+-----
xb
  x1cat |
    0-  |           0 (base)
```



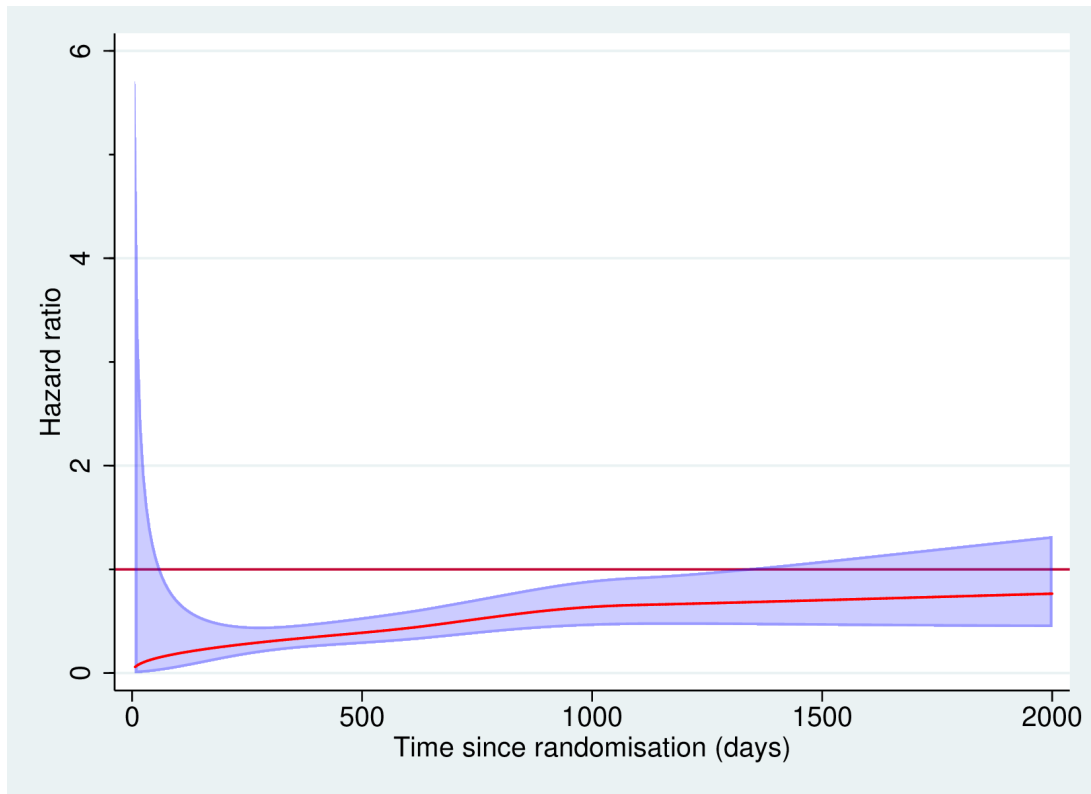


Figure 4: Time-varying hazard ratio for progesterone receptor  $\geq 20$  fmol, German Breast Cancer Study Group

### Q3

- (a) For a cohort study investigating the onset of diabetes, assume that we have linked general practice (GP) visits with a diabetes quality register and the population register. For each individual diagnosed with diabetes, we assume that diabetes onset happens between the last GP visit when an individual was not diagnosed and the first visit when the individual was diagnosed with diabetes. Also assume that individuals enter the cohort study at different ages and are followed through to death, emigration or 2016-12-01, whichever happens first. Discuss this study design in terms of truncation and censoring. (2pts)
- (b) Consider a cancer patient cohort study with two groups defined by cancer stage at diagnosis. The first group has localised or regional spread at cancer diagnosis, and the second group has metastatic spread at cancer diagnosis. For the first group, the five-year survival is 0.8. For the second group, assume we have proportional hazards with a hazard ratio of 2. What is the five-year survival in the second group? Show your working. (1pt)
- (c) Compare and contrast (i) a cohort study analysed using Cox regression and (ii) a nested case-control study analysed using conditional logistic regression. How are these two designs and analyses related? What are the advantages and disadvantages of each? (3pts)